# Stat 515: Introduction to Statistics

Chapter 10

- The **response variable** is the variable of interest to be measured in an experiment
  - Essentially this is what we would like to model or predict
- Factors are the variable whose effect on the response is of interest
  - Essentially this is what we would like to use to model or predict the **response**

 Note: Both the response(s) and factor(s) are either qualitative, quantitative discrete or quantitative continuous just as we've learned about variables earlier in the semester.

- Factor Levels are the values of the factor
  - We most often refer to the categories that make up a qualitative variable as 'levels'
- **Treatments** of an experiment are the factor level combinations utilized

- An experimental unit is the object on which the response and factors are observed
  - These are often, but not always people

- Recall the difference between designed experiments and observational studies
- An observational study measures the response variable without attempting to influence the value of either the response or explanatory variables.
- A **designed study** occurs when a researcher assigns the individuals or subjects into groups and intentionally affects their explanatory variables (think treatments)

- C. Myrray Parkes headed a study of 4,486 men of 55 years of age and older who had their wives die in 1957. For up to nine years, these widowers were tracked and 213 died during the first six months – that's about 5%..
- This experiment is a observational study. C Myrray Parkes didn't murder 4,486 women in 1957 just to do this study.

- Herbet Benson, MD headed a study in 2005 to see if intercessory prayer influenced recovery from bypass surgery. There were three groups in the study: 1. Those being prayed for that didn't know 2. Those being prayed for that did know 3. Those not being prayed for
- This is a designed study because the researchers assigned different patients to different groups; they controlled who was prayed for and who wasn't instead of just observing and asking the families whether or not they had friends and families praying for the patient.

 Note: both of these studies were planned but we only poke and prod the experimental units in the second study where the researched assigned different "treatments" to each patient.

- A completely randomized designed is a design where the treatments are randomly assigned to the experimental units
- Note: Why we need a random sample is obvious by now from the need in Ch 7-9 to have our inference be applicable to the rest of the population – why the treatments need to be random is less obvious

 We call an experiment balanced if we assign the same number of experimental units to each treatment

• Consider a field of plots where we plant corn



• Here, we consider two different fertilizers and we want to know which works the best



 Consider this very 'un-random' assortment of fertilizers A & B, what's a possible drawback?



• Perhaps, the farm house puts shade on the four plots to the right. In this case B would grow less, perhaps even if B is superior to A.



 Consider a field with poor drainage suppose that the plots to the right puddled with water after rain. In this case B would grow less, perhaps even if B is superior to A



• In either case, randomization would have helped us. Something like the randomization below would put A and B in both situations.



• In either case, randomization would have helped us. Something like the randomization below would put A and B in both situations.



• In either case, randomization would have helped us. Something like the randomization below would put A and B in both situations.



# Single Factor Design

 In a Single Factor Design we are only considering one factor in the modeling/prediction of the response variable

• We're interested in testing

 $H_0: \mu_1 = \mu_2 = \cdots = \mu_k$ : The means are equal

 $H_A$ : At least one of the  $\mu_i$  is different

• Sum of Squares for Treatments (SST): the variation between the treatment means

$$SST = \sum n_i (\bar{x_i} - \bar{x})^2$$

 $n_i$  = sample size for treatment *i*  $\overline{x_i}$  = sample mean for treatment i  $\overline{x}$  = sample mean overall

• Sum of Squares for Error (SSE): the variation within the treatments

$$SSE = \sum (x_{1j} - \overline{x_1})^2 + \sum (x_{2j} - \overline{x_2})^2 + \dots + \sum (x_{kj} - \overline{x_k})^2$$
$$x_{ij} = observations \ j \ from \ treatment \ 1$$
$$\overline{x_i} = sample \ mean \ from \ treatment \ i$$
$$\overline{x} = sample \ mean \ overall$$

• Sum of Squares for Error (SSE): the variation within the treatments

**Recalling** 
$$s^2$$
:  
 $SSE = (n_1 - 1)s_1^2 + (n_2 - 1)s_2^2 + \dots + (n_k - 1)s_k^2$ 

 $n_i$  = sample size for treatment *i*  $s_i^2$  = sample variance for treatment i

 Mean Square for Treatments (MST): measures the variation among the treatment

$$MST = \frac{SST}{k-1}$$
  
SST = sum of squares for treatments  
k = the number of treatments

• Mean Square for Error (MSE): measures the variation among the treatment

$$MSE = \frac{SSE}{n-k}$$
  

$$SSE = sum of squares for Error$$
  

$$n = the number of experimental units$$
  

$$k = the number of treatments$$

 F Statistic(F): Completes the test that we're interested in

$$F = \frac{MST}{MSE}$$
  

$$SSE = sum of squares for Error$$
  

$$n = the number of experimental units$$
  

$$k = the number of treatments$$

 $H_0: \mu_1 = \mu_2 = \dots = \mu_k$ : The means are equal  $H_A$ : At least one of the  $\mu_i$  is different Assumptions: 1) Samples are randomly selected 2) The k distributions are normal 3)  $\sigma_1 = \sigma_2 = \cdots = \sigma_k$ **Test Statistic:**  $F = \frac{MST}{MSE}$ **Reject when:**  $F > F_{1-\alpha} = qf(1-\alpha, k-1, n-k)$ 

• Suppose we have the following results for corn yield (in bushels) per plot of land with either fertilizer A or fertilizer B



- $\overline{x_A} = 99.25$
- $\overline{x_B} = 103$
- $s_A^2 = 9.58333$
- $s_B^2 = 10$

 In this case we're only considering one factor: fertilizer which has two levels A and B.

• We're interested in testing  $H_0: \mu_A = \mu_B:$  The fertilizer work equally  $H_A: \mu_A \neq \mu_B:$  The fertilizer don't work equally

 $SST = \sum n_i (\bar{x}_i - \bar{x})^2$ = 4(99.25 - 101.125)<sup>2</sup> + 4(103 - 101.125)<sup>2</sup> = 28.125

$$SSE = (n_A - 1)s_A^2 + (n_B - 1)s_B^2$$
  
= (4 - 1)9.58333 + (4 - 1)10  
= 58.7499

$$MST = \frac{SST}{k-1} = \frac{28.125}{2-1} = 28.125$$

$$MSE = \frac{58.7499}{8-2} = 9.79165$$

 $F = \frac{MST}{MSE} = \frac{28.125}{9.79165} = 2.872345$ 

**Reject when:**  $F > F_{1-\alpha} = qf(1 - \alpha, k - 1, n - k)$ *F*=2.872345

 $F_{1-\alpha} = qf(.95, 2-1, 8-2) = 5.987378$ 

**2.872345<5.987378** so we reject the null hypothesis in favor of the alternative.

**Note:** our pvalue agrees with this decision

## In R

- plotYield<-c(102,101,95,99,102,99,106,105)
  fert<-c('A','A','A','A','B','B','B','B')
  farming<-lm(plotYield~fert)
  summary(farming)
  anova(farming)</pre>
- This gives us the F statistic and the associated pvalue, along with other measurements

## In R

#### summary(farming)

```
Call:
lm(formula = plotYield ~ fert)
```

#### Residuals:

Min 1Q Median 3Q Max -4.250 -1.750 0.750 2.188 3.000

#### Coefficients:

	Estimate Std.	Error	t value	Pr(> t )	
(Intercept)	99.250	1.565	63.435	1.03e-09	***
fertB	3.750	2.213	1.695	0.141	

Signif. codes: 0 `\*\*\*' 0.001 `\*\*' 0.01 `\*' 0.05 `.' 0.1 ` ' 1

Residual standard error: 3.129 on 6 degrees of freedom Multiple R-squared: 0.3237, Adjusted R-squared: 0.211 F-statistic: 2.872 on 1 and 6 DF, p-value: 0.141

#### Mean Difference for B t test for significance F statistic

## In R

#### anova(farming)

